STA 360/602L: MODULE 2.6

LOSS FUNCTIONS AND BAYES RISK

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BAYES ESTIMATE

- As we've seen by now, having posterior distributions instead of onenumber summaries is great for capturing uncertainty.
- That said, it is still very appealing to have simple summaries, especially when dealing with clients or collaborators from other fields, who desire one.
- Can we obtain a single estimate of θ based on the posterior? Sure!
- **Bayes estimate** is the value $\hat{\theta}$, that minimizes the Bayes risk.



BAYES ESTIMATE

- Bayes risk is defined as the expected loss averaged over the posterior distribution.
- Put differently, a Bayes estimate $\hat{\theta}$ has the lowest posterior expected loss.
- That's fine, but what does expected loss mean?
- Frequentist risk also exists but we won't go into that here.



LOSS FUNCTIONS

- A loss function $L(\theta, \delta(y))$ is a function of a parameter θ , where $\delta(y)$ is some decision about θ , based on just the data y.
- For example, $\delta(y) = \bar{y}$ can be the decision to use the sample mean to estimate θ , the true population mean.
- $L(\theta, \delta(y))$ determines the penalty for making the decision $\delta(y)$, if θ is the true parameter; $L(\theta, \delta(y))$ characterizes the price paid for errors.



LOSS FUNCTIONS

 A common choice for example, when dealing with point estimation, is the squared error loss, which has

$$L(\theta, \delta(y)) = (\theta - \delta(y))^2.$$

Bayes risk is thus

$$ho(heta,\delta) = \mathbb{E}\left[\left[\left. L(heta,\delta(y)) \left| y
ight] = \int L(heta,\delta(y)) \cdot \pi(heta|y) \; d heta,
ight.$$

and we proceed to find the value $\hat{\theta},$ that is, the decision $\delta(y),$ that minimizes the Bayes risk.



BAYES ESTIMATOR UNDER SQUARED ERROR LOSS

- Turns out that, under squared error loss, the decision $\delta(y)$ that minimizes the posterior risk is the posterior mean.
- Proof: Let $L(heta,\delta(y))=(heta-\delta(y))^2.$ Then,

$$egin{aligned} &
ho(heta,\delta) = \int L(heta,\delta(y))\cdot\pi(heta|y)\;d heta.\ &= \int (heta-\delta(y))^2\cdot\pi(heta|y)\;d heta. \end{aligned}$$

- Expand, then take the partial derivative of $\rho(\theta, \delta)$ with respect to $\delta(y)$.
- To be continued on the board!



BAYES ESTIMATOR UNDER SQUARED ERROR LOSS

$$p(heta,\delta)\int (heta-\delta(y))^2\cdot\pi(heta|y)\;d heta.$$

- Easy to see then that $\delta(y) = \mathbb{E}[heta|x]$ is the minimizer.
- Well that's great! The posterior mean is often very easy to calculate in most cases.
- In the beta-binomial case for example, the Bayes estimate under squared error loss is just

$$\hat{ heta} = rac{a+y}{a+b+n},$$

the posterior mean.



WHAT ABOUT OTHER LOSS FUNCTIONS?

 Clearly, squared error is only one possible loss function. An alternative is absolute loss, which has

 $L(heta,\delta(y))=| heta-\delta(y)|.$

- Absolute loss places less of a penalty on large deviations & the resulting Bayes estimate is posterior median.
- Median is actually relatively easy to estimate.



WHAT ABOUT OTHER LOSS FUNCTIONS?

• Recall that for a continuous random variable Y with cdf F, the median of the distribution is the value z, which satisfies

$$F(z)=\Pr(Y\leq z)=rac{1}{2}=\Pr(Y\geq z)=1-F(z).$$

- As long as we know how to evaluate the CDF of the distribution we have, we can solve for z.
- Think R!



WHAT ABOUT OTHER LOSS FUNCTIONS?

 For the beta-binomial model, the CDF of the beta posterior can be written as

$$F(z)=\Pr(heta\leq z|y)=\int_0^z ext{Beta}(heta|a+y,b+n-y)d heta.$$

- Then, if $\hat{ heta}$ is the median, we have that $F(\hat{ heta})=0.5.$
- To solve for $\hat{ heta}$, apply the inverse CDF $\hat{ heta}=F^{-1}(0.5).$
- In R, that's simply

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qbeta(0.5,a+y,b+n-y)
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• For other popular distributions, switch out the beta.



LOSS FUNCTIONS AND DECISIONS

- Loss functions are not specific to estimation problems but are a critical part of decision making.
- For example, suppose you are deciding how much money to bet (\$A) on Duke in the next UNC-Duke men's basketball game.
- Suppose, if Duke
 - loses (y = 0), you lose the amount you bet (\$A)
 - wins (y = 1), you gain B per \$1 bet
- What is a good sampling distribution for y here?
- Then, the loss function can be characterized as

L(A, y) = A(1 - y) - y(BA),

with your action being the amount bet A.

When will your bet be "rational"?



y is an unknown state, but we can think of it as a new prediction y_{n+1} given that we have data from win-loss records (y_{1:n}) that can be converted into a Bayesian posterior,

 $heta \sim ext{beta}(a_n, b_n),$

with this posterior concentrated slightly to the left of 0.5, if we only use data on UNC-Duke games (UNC men lead Duke 139-112 all time).

- Actually, it might make more sense to focus on more recent head-to-head data and not the all time record.
- In fact, we might want to build a model that predicts the outcome of the game using historical data & predictors (current team rankings, injuries, etc).
- However, to keep it simple for this illustration, go with the posterior above.



• The Bayes risk for action A is then the expectation of the loss function,

 $ho(A) = \mathbb{E}\left[\left| L(A,y) \right| y_{1:n}
ight].$

- To calculate this as a function of A and find the optimal A, we need to marginalize over the **posterior predictive distribution** for y.
- Why are we using the posterior predictive distribution here instead of the posterior distribution?
- As an aside, recall from Module 2.3 that

• By the way, what do a_n and b_n represent?

$$p(y_{n+1}|y_{1:n})=rac{a_n^{y_{n+1}}b_n^{1-y_{n+1}}}{a_n+b_n}; \hspace{0.2cm} y_{n+1}=0,1.$$

• Specifically, that the posterior predictive distribution here is $\operatorname{Bernoulli}(\hat{\theta})$, with

$$\hat{ heta} = rac{a_n}{a_n+b_n}$$



• With the loss function L(A, y) = A(1 - y) - y(BA), and using the notation y_{n+1} instead of y (to make it obvious the game has not been played), the Bayes risk (expected loss) for bet A is

$$egin{aligned} &
ho(A) = \mathbb{E}\left[\left.L(A,y_{n+1})\mid y_{1:n}
ight] \ &= \mathbb{E}\left[A(1-y_{n+1})-y_{n+1}(BA)\mid y_{1:n}
ight] \ &= A \ \mathbb{E}\left[1-y_{n+1}\mid y_{1:n}
ight] - (BA) \ \mathbb{E}\left[y_{n+1}\mid y_{1:n}
ight] \ &= A \ (1-\mathbb{E}\left[y_{n+1}\mid y_{1:n}
ight]) - (BA) \ \mathbb{E}\left[y_{n+1}\mid y_{1:n}
ight] \ &= A \ (1-\mathbb{E}\left[y_{n+1}\mid y_{1:n}
ight]) - (BA) \ \mathbb{E}\left[y_{n+1}\mid y_{1:n}
ight] \ &= A \ (1-\mathbb{E}\left[y_{n+1}\mid y_{1:n}
ight] \ (1+B)) \,. \end{aligned}$$



Hence, your bet is rational as long as

$$\mathbb{E}\left[y_{n+1} | \; y_{1:n}
ight](1+B) > 1
onumber \ rac{a_n(1+B)}{a_n+b_n} > 1.$$

- Clearly, there is no limit to the amount you should bet if this condition is satisfied (the loss function is clearly too simple).
- Loss function needs to be carefully chosen to lead to a good decision finite resources, diminishing returns, limits on donations, etc.
- Want more on loss functions, expected loss/utility, or decision problems in general? Consider taking a course on decision theory (STA623?).



WHAT'S NEXT?

Move on to the readings for the next module!

